

Nonparametric Estimation of the Cumulative Intensity Function for a Nonhomogeneous Poisson Process from Multiple Realizations with Nonoverlapping and Overlapping Intervals

Larry Leemis

Department of Mathematics

The College of William & Mary

Williamsburg, VA 23187-8795 USA

E-mail: leemis@math.wm.edu

CSCI 420 / Math 410 UMSA Research Seminar

February 7, 2002

Outline

1. Motivation
2. Probabilistic properties
3. Estimating $\Lambda(t)$ from k realizations on $(0, S]$
4. Estimating $\Lambda(t)$ from overlapping realizations
5. Conclusions

Note: Some of this work is with with Brad Arkin (RST Corporation). This work was supported by an NSF grant supporting the UMSA (Undergraduate Modeling and Simulation Analysis) REU (Research Experience for Undergraduates) at The College of William & Mary.

1. Motivation

Although easy to estimate and simulate, HPPs and renewal processes do not allow for varying rates. The use of an NHPP is often more appropriate for modeling.

Example: Customer arrivals to a fast food restaurant



Other examples:

Cyclone arrival times in the Arctic Sea (Lee, Wilson, and Crawford, 1991)

Database transaction times (Lewis and Shedler, 1976)

Calls for on-line analysis of electrocardiograms at a hospital in Houston (Kao and Chang, 1988)

Respiratory cancer deaths near a steel complex in Scotland (Lawson, 1988)

Repairable systems: blood analyzers, fan motors, power supplies, turbines (Nelson, 1988)

2. Probabilistic properties

Notation

t	time
$N(t)$	number of events by time t
$\lambda(t)$	instantaneous arrival rate at time t (intensity function)
$\Lambda(t) = \int_0^t \lambda(\tau) d\tau$	cumulative intensity function

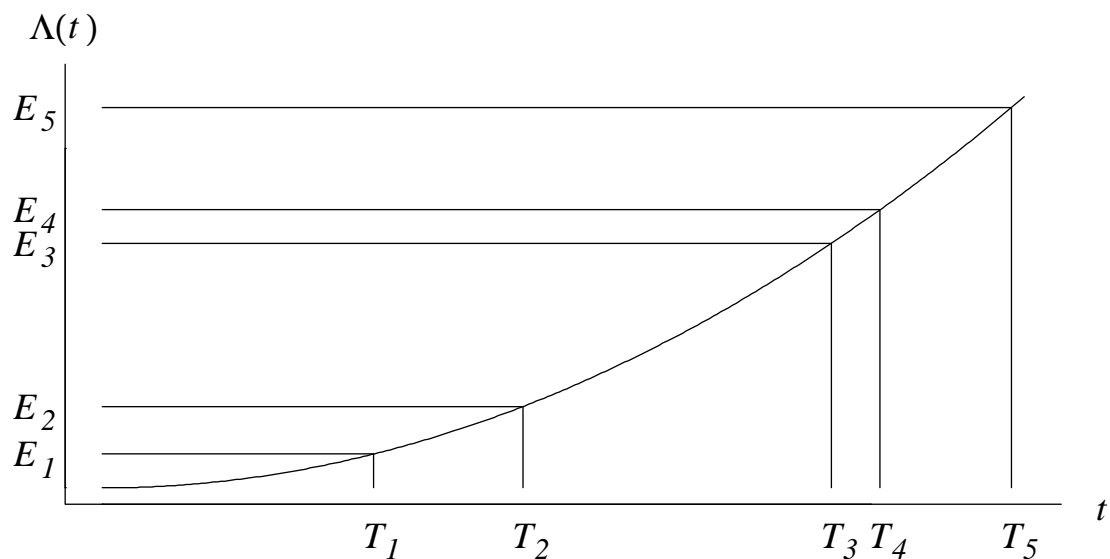
Property 1

$$\Pr(N(b) - N(a) = n) = \frac{\left[\int_a^b \lambda(\tau) d\tau \right]^n e^{-\int_a^b \lambda(\tau) d\tau}}{n!} \quad n = 0, 1, \dots$$

Property 2

$$E[N(t)] = \Lambda(t)$$

Property 3 (Çinlar, 1975) If E_1, E_2, \dots are event times in a *unit* HPP then $\Lambda^{-1}(E_1), \Lambda^{-1}(E_2), \dots$ are event times in an NHPP with cumulative intensity function $\Lambda(t)$.



3. Estimating $\Lambda(t)$ from k realizations on $(0, S]$

Data

t	time
$(0, S]$	time interval where observations are collected
k	number of realizations collected
n_1, n_2, \dots, n_k	number of observations per realization
$t_{(1)}, t_{(2)}, \dots, t_{(n)}$	superposition of observations
$n = \sum_{i=1}^k n_i$	total number of observations collected

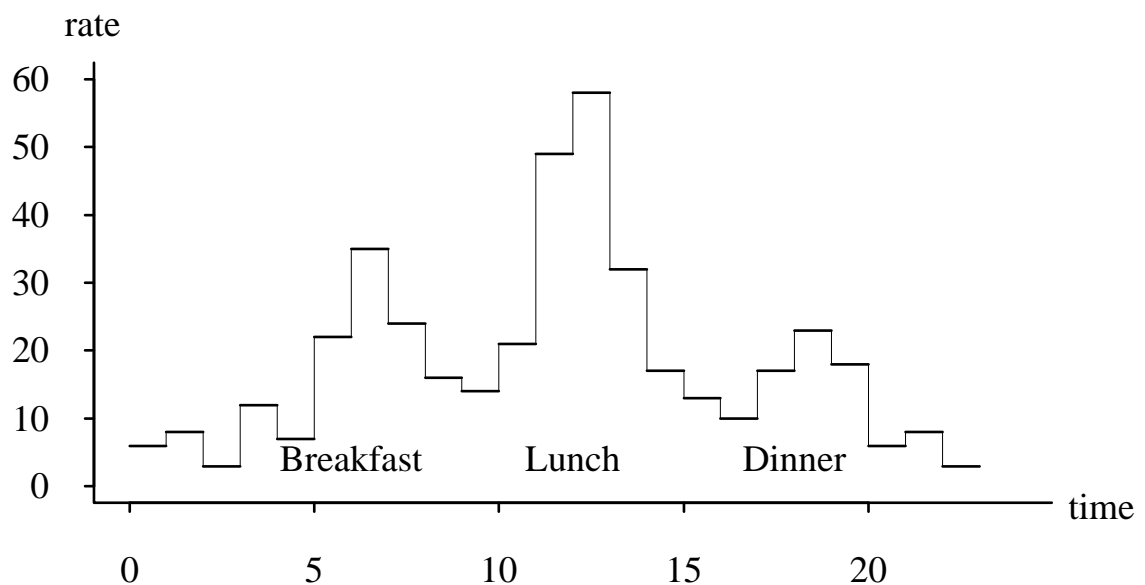
Intuitive solution (Law and Kelton, 2000): partition time axis and let $\lambda(t)$ be piecewise constant.

Problems

(a) Determining cell width

- Small cell width — sampling variability
- Large cell width — miss trend

(b) Subjective due to arbitrary parameters

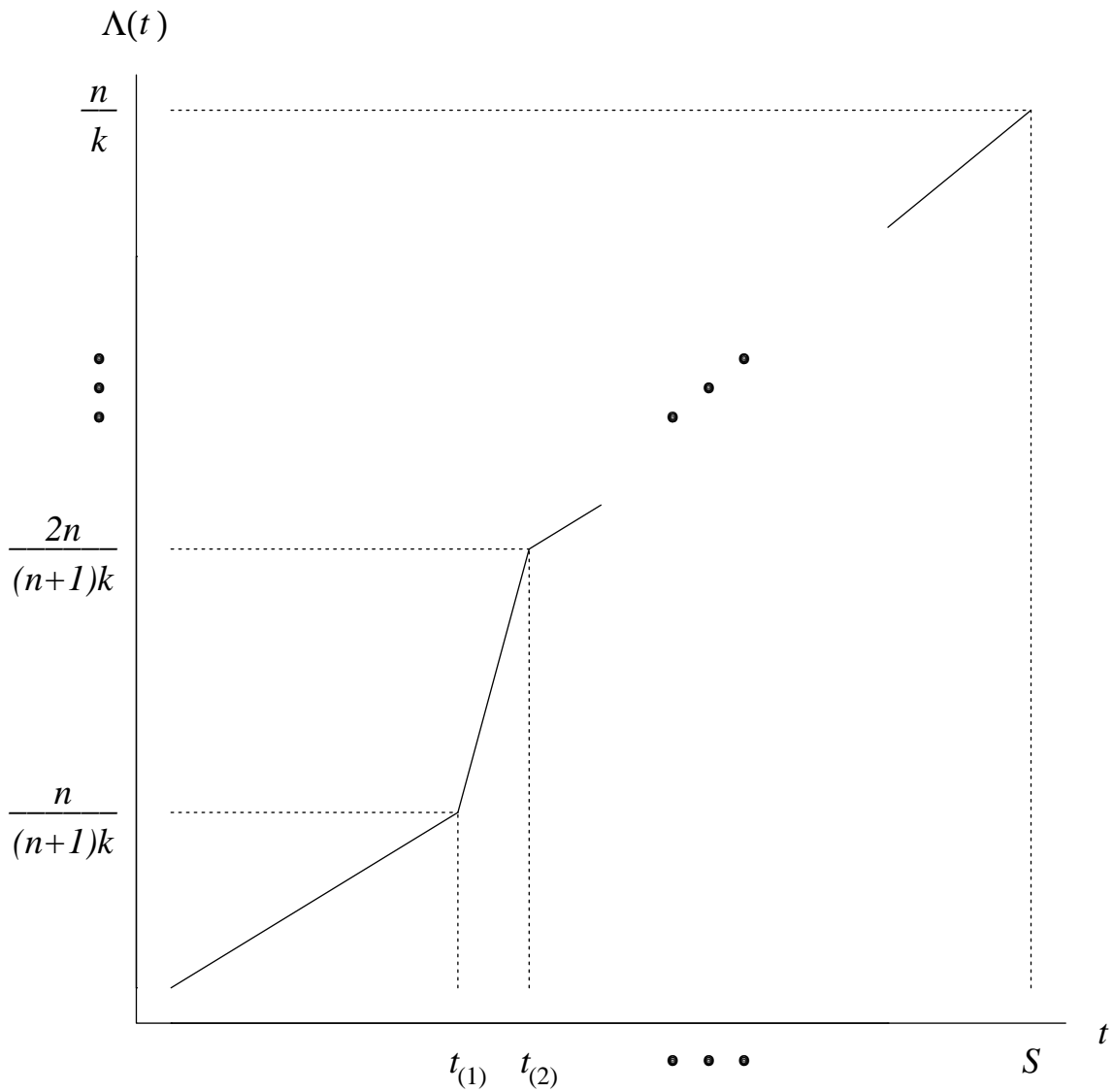


Piecewise-linear nonparametric cumulative intensity function estimator

$$\hat{\Lambda}(t) = \frac{in}{(n+1)k} + \left[\frac{n(t - t_{(i)})}{(n+1)k(t_{(i+1)} - t_{(i)})} \right] \quad t_{(i)} < t \leq t_{(i+1)}$$

for $i = 0, 1, 2, \dots, n$.

Rationale: $\hat{\Lambda}(S) = n/k$



Properties

- Handles ties as expected
- Consistency

$$\lim_{k \rightarrow \infty} \hat{\Lambda}(t) = \Lambda(t)$$

- Confidence interval (asymptotically exact) for $\Lambda(t)$

$$\hat{\Lambda}(t) \pm z_{\alpha/2} \sqrt{\frac{\hat{\Lambda}(t)}{k}}$$

- Variate generation straightforward

Input:

Number of observed arrivals n

Number of active realizations k

Superpositioned observations $t_{(1)}, t_{(2)}, \dots, t_{(n+1)}$

Output:

Event times T_1, T_2, \dots, T_{i-1} on $(0, S]$

```
 $i \leftarrow 1$  [initialize variate counter]
generate  $U_i \sim U(0, 1)$  [generate initial random number]
 $E_i \leftarrow -\log_e(1 - U_i)$  [generate initial exponential variate]
while  $E_i < n/k$  do
  begin
     $m \leftarrow \lfloor \frac{(n+1)kE_i}{n} \rfloor$  [set  $m \ni \hat{\Lambda}(t_{(m)}) < E_i \leq \hat{\Lambda}(t_{(m+1)})$ ]
     $T_i \leftarrow t_{(m)} + [t_{(m+1)} - t_{(m)}] \left( \frac{(n+1)kE_i}{n} - m \right)$  [generate event time]
   $i \leftarrow i + 1$  [increment variate counter]
  generate  $U_i \sim U(0, 1)$  [generate next random number]
   $E_i \leftarrow E_{i-1} - \log_e(1 - U_i)$  [generate next HPP event time]
end
```

4. Estimating $\Lambda(t)$ from overlapping realizations

Data

t time
 $(0, S]$ time interval where observations are collected
 r # time intervals where the # realizations is fixed
 $(s_j, s_{j+1}]$ interval $j + 1$, $j = 0, 1, \dots, r - 1$
 k_{j+1} # realizations observed on $(s_j, s_{j+1}]$, $j = 0, 1, \dots, r - 1$
 n_{j+1} number of observations on $(s_j, s_{j+1}]$
 $t_{(0)}, t_{(1)}, \dots, t_{(n+r)}$ superposition of observations, s_0, s_1, \dots, s_r
Note: $s_0 = 0, s_r = S$

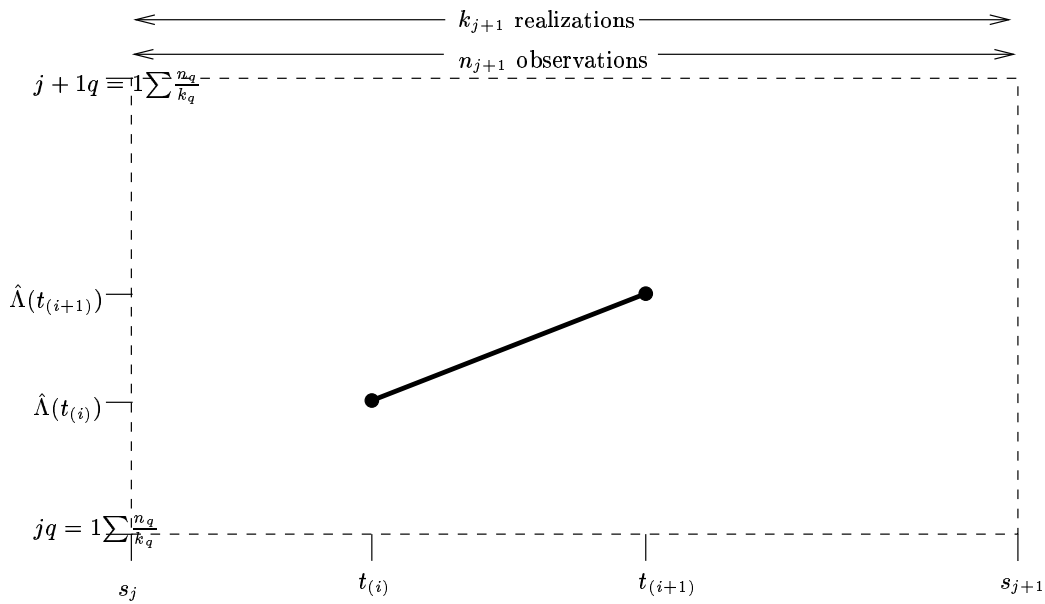
$$\hat{\Lambda}(t) = \sum_{q=1}^j \frac{n_q}{k_q} + \frac{(i - \sum_{q=1}^j (n_q + 1)) n_{j+1}}{(n_{j+1} + 1) k_{j+1}} + \left[\frac{n_{j+1} (t - t_{(i)})}{(n_{j+1} + 1) k_{j+1} (t_{(i+1)} - t_{(i)})} \right],$$

$$t_{(i)} < t \leq t_{(i+1)}; \quad i = 0, 1, 2, \dots, n + r - 1,$$

$$s_j < t \leq s_{j+1}; \quad j = 0, 1, \dots, r - 1,$$

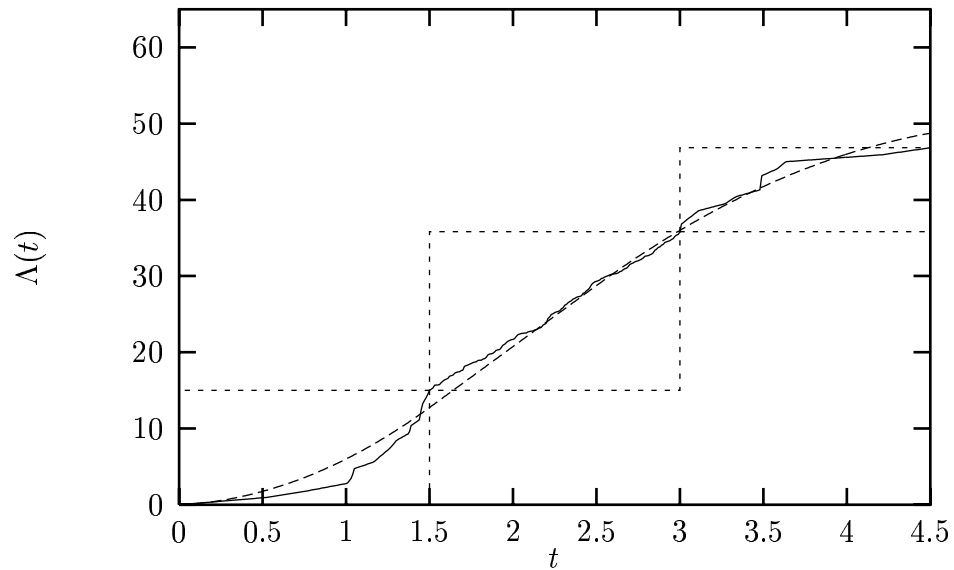
Rationale: $\hat{\Lambda}(s_{j+1}) = \sum_{q=1}^{j+1} n_q / k_q$

Single segment of $\hat{\Lambda}(t)$ in the $(j + 1)$ st region:



Example: Lunchwagon arrivals (Klein and Roberts, 1984)

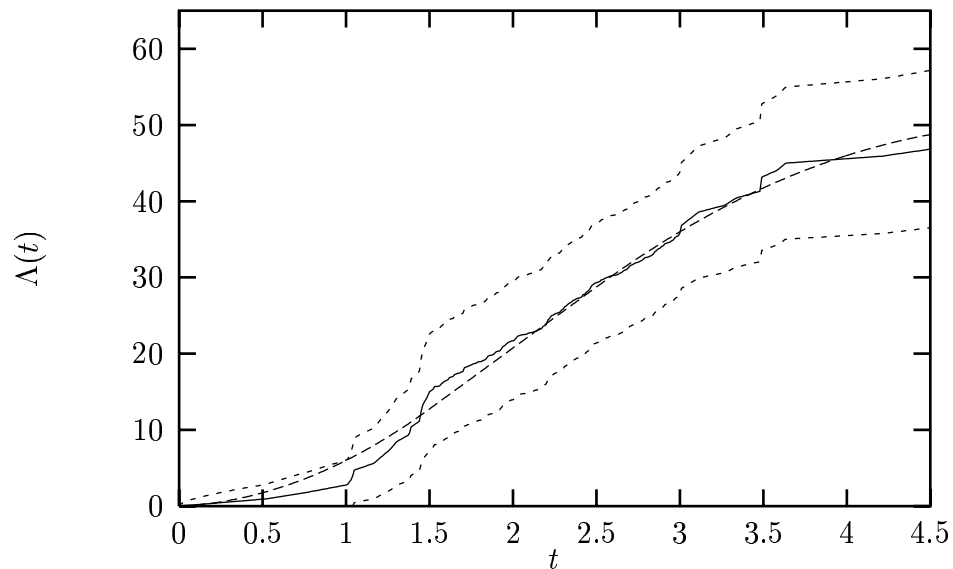
Depiction of three regions for lunchwagon arrivals from 10:00 AM to 2:30 PM for $k_1 = 1$, $k_2 = 12$, and $k_3 = 1$; $s_1 = 1.5$, $s_2 = 3$, and $s_3 = 4.5$.



An asymptotically exact $100(1 - \alpha)\%$ confidence interval for $\Lambda(t)$

$$|\Lambda(t) - \hat{\Lambda}(t)| < z_{\alpha/2} \sqrt{\frac{\hat{\Lambda}(t) - \hat{\Lambda}(s_j)}{k_{j+1}} + \sum_{q=1}^j \frac{\hat{\Lambda}(s_q) - \hat{\Lambda}(s_{q-1})}{k_q}}$$

Parent cumulative intensity function, nonparametric estimator, and 95% confidence bands for lunchwagon arrivals from 10:00 AM to 2:30 PM for $k_1 = 1$, $k_2 = 12$, and $k_3 = 1$; $s_1 = 1.5$, $s_2 = 3$, and $s_3 = 4.5$.



Variate Generation

Input:

Number of partitions r

Number of active realizations k_1, k_2, \dots, k_r

Number of observed arrivals per partition n_1, n_2, \dots, n_r

Superpositioned values $t_{(0)}, t_{(1)}, \dots, t_{(n+r)}$

Output:

Event times T_1, T_2, \dots, T_{i-1} on $(0, S]$

```
 $i \leftarrow 1$  [initialize variate counter]
 $j \leftarrow 0$  [initialize region counter]
 $\text{MAX} \leftarrow \sum_{q=1}^r n_q/k_q$  [set MAX to  $\hat{\Lambda}(S)$ ]
generate  $U_i \sim U(0, 1)$  [generate initial random number]
 $E_i \leftarrow -\log_e(1 - U_i)$  [generate initial exponential variate]
while  $E_i < \text{MAX}$  do
  begin
    while  $E_i > \sum_{q=1}^{j+1} n_q/k_q$  do [update  $j$  if necessary]
      begin
         $j \leftarrow j + 1$  [increment region counter]
      end
       $m \leftarrow \left\lfloor \frac{(n_{j+1}+1)k_{j+1}(E_i - \sum_{q=1}^j n_q/k_q)}{n_{j+1}} \right\rfloor + \sum_{q=1}^j (n_q + 1)$ 
      [set  $m \ni \hat{\Lambda}(t_{(m)}) < E_i \leq \hat{\Lambda}(t_{(m+1)})$ ]
       $T_i \leftarrow t_{(m)} + [t_{(m+1)} - t_{(m)}] \left( \frac{(n_{j+1}+1)k_{j+1}(E_i - \sum_{q=1}^j n_q/k_q)}{n_{j+1}} - (m - \sum_{q=1}^j (n_q + 1)) \right)$ 
      [generate event time]
     $i \leftarrow i + 1$  [increment variate counter]
    generate  $U_i \sim U(0, 1)$  [generate next random number]
     $E_i \leftarrow E_{i-1} - \log_e(1 - U_i)$  [generate next HPP event time]
  end
```

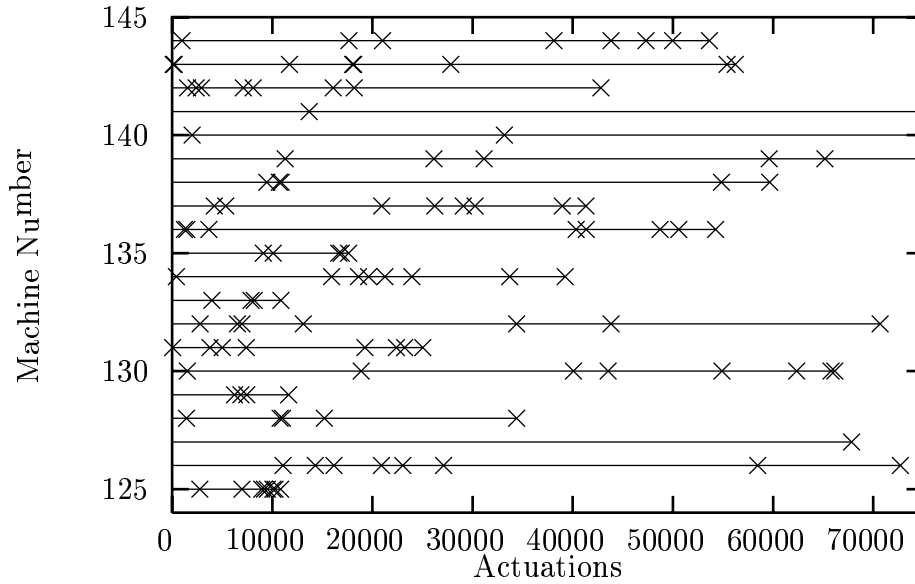
Example 1: Monte Carlo evaluation of the confidence interval for $\Lambda(t)$

Coverages in the lunchwagon example (nominal coverage 0.95; 100,000 replications; $k_1 = 1, k_2 = 12, k_3 = 1$; $s_0 = 0, s_1 = 1.5, s_2 = 3, s_3 = 4.5$).

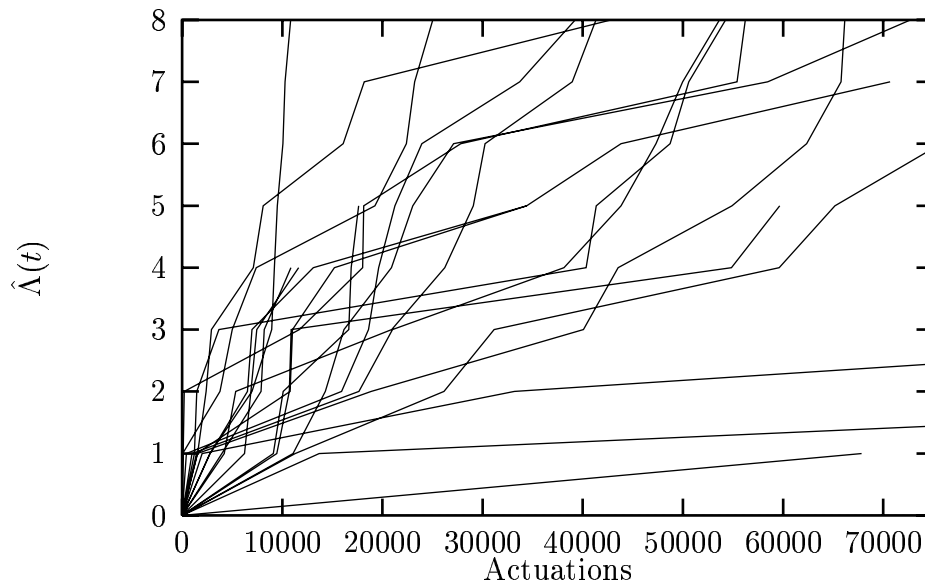
Time	Actual Coverage	Misses High	Misses Low
0.90	0.9501	0.0013	0.0487
1.35	0.9386	0.0048	0.0566
1.80	0.9505	0.0200	0.0296
2.25	0.9466	0.0196	0.0339
2.70	0.9498	0.0174	0.0329
3.15	0.9509	0.0295	0.0196
3.60	0.9498	0.0251	0.0251
4.05	0.9517	0.0167	0.0316

Example 2: Failure times for 20 copy machines (Zaino and Berke 1992)

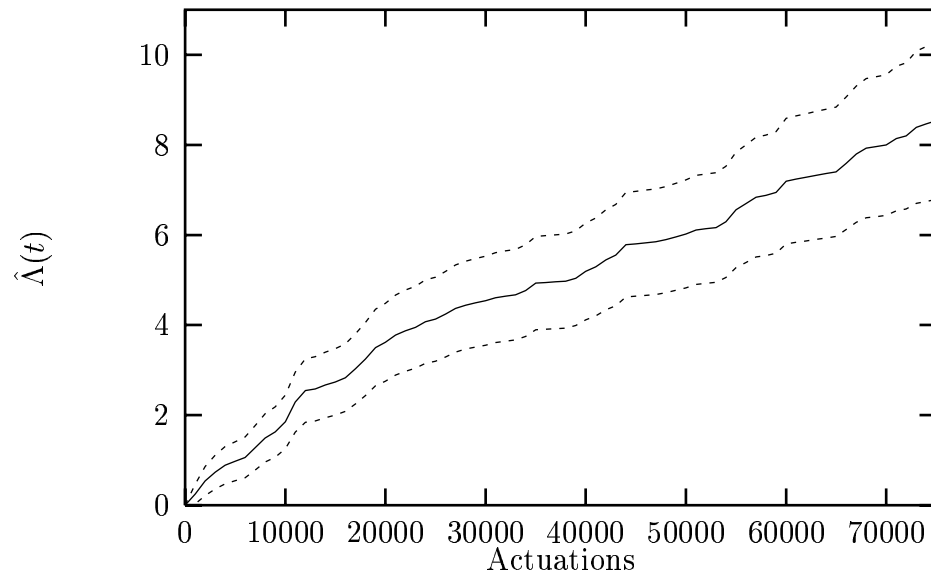
Failure times



$\hat{\Lambda}(t)$ for each machine



$\hat{\Lambda}(t)$ for the copy machine failure times

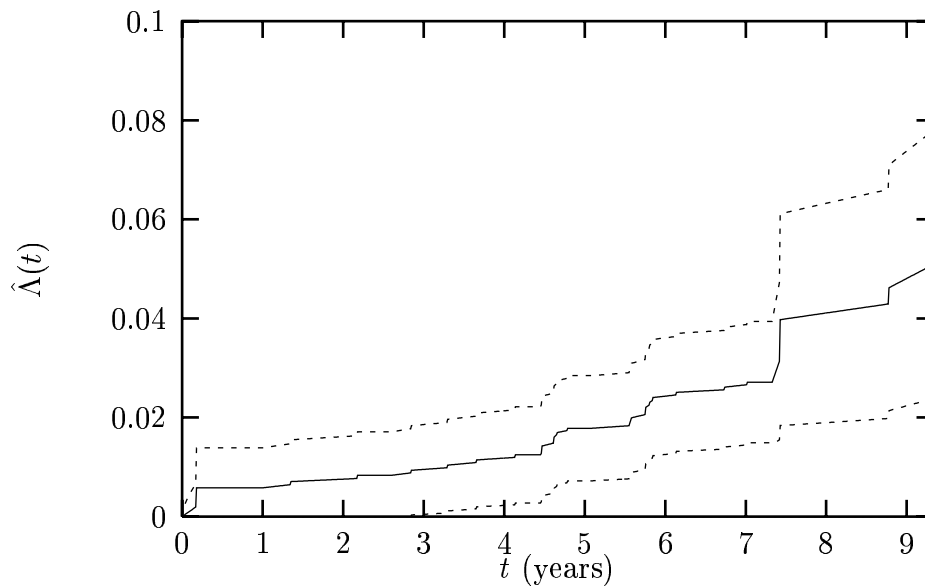


Example 3: Failure times of heat pump compressors (Nelson 1990)

The compressors are located in five separate buildings, each under repair contract for a time span $(a, b]$, indicated by the bold-face values. The data set consists of $n = 28$ failure times, and yields $r = 29$ regions.

Bldg	Num of Comp	Entry time, Failure Times, Exit time
B	164	2.59 , 3.30, 4.62, 4.62, 5.75, 5.75, 7.42, 7.42, 8.77, 9.27, 9.27, 9.33
D	356	4.45 , 4.47, 4.47, 5.56, 5.57, 5.80, 6.13, 7.02, 7.05
E	458	1.00 , 2.85, 4.65, 4.79, 5.85, 6.73, 7.33
H	149	0.00 , 0.17, 0.17, 1.34, 5.09
K	195	0.00 , 2.17, 3.65, 4.14, 4.14

$\hat{\Lambda}(t)$ for the heat pump compressor failure times



5. Conclusions

- (a) Nonparametric estimation and simulation for NHPPs is straightforward
- (b) Collecting data across overlapping intervals does not pose any significant problems
- (c) Once coded, this approach requires less effort than a parametric renewal process in order to simulate the observations